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UNIVERSITY OF PRIMORSKA

FAMNIT, MATHEMATICS

PROBABILITY

MIDTERM 2

JUNE 4th, 2021

INSTRUCTIONS

Read carefully the text of the problems before attempting to solve them. Five problems out of six count for 100%. You are allowed one A4 sheet with formulae and theorems, and a handbook of mathematics. Time allowed: 120 minutes.

Question	a.	b.	c.	d.	Total
1.			•	•	
2.			•	•	
3.			•	•	
4.			•	•	
5.			•	•	
6.			•	•	
Total					

1. (20) An urn contains B black and R red balls. We select balls at random without replacement until we pick the first red ball at which point we stop. Let X be the number of black balls selected.

a. (10) Compute $E(X)$.

Hint: use indicators.

Solution: number all the black balls with $k = 1, 2, \dots, B$. Letting

$$I_k = \begin{cases} 1 & \text{if the } k\text{-th black ball comes before the first red ball} \\ 0 & \text{else,} \end{cases}$$

we have

$$X = \sum_{k=1}^B I_k.$$

By symmetry, all indicators have the same expectation. To compute the probability $P(I_1 = 1)$, note that only the relative positions of the k -th black ball and the R red balls are relevant. Since all the permutations are equally likely we get

$$P(I_1 = 1) = \frac{1}{R+1},$$

and as a consequence

$$E(X) = \frac{B}{R+1}.$$

b. (10) Compute $\text{var}(X)$.

Solution: we use the indicators from the first part. First observe that

$$\text{var}(I_k) = \frac{R}{(R+1)^2}$$

for all k . Again by symmetry, all the covariances $\text{cov}(I_k, I_l)$ for $k \neq l$ are also equal. We find $P(I_k = 1, I_l = 1)$ by considering relative positions of the k -th and l -th black ball, and the R red balls. The probability that the two black balls are the first two is

$$P(I_k = 1, I_l = 1) = 2 \cdot \frac{1}{R+1} \cdot \frac{1}{R+2}.$$

It follows,

$$\text{cov}(I_k, I_l) = \frac{2}{(R+1)(R+2)} - \frac{1}{(R+1)^2},$$

which simplifies to

$$\text{cov}(I_k, I_l) = \frac{R}{(R+1)^2(R+2)}.$$

By the formula for the variance of sums we have

$$\text{var}(X) = B \cdot \frac{R}{(R+1)^2} + B(B-1) \cdot \frac{R}{(R+1)^2(R+2)},$$

which simplifies to

$$\text{var}(X) = \frac{BR(B+R+1)}{(R+1)^2(R+2)}.$$

2. (20) Let X_1, X_2, \dots be independent geometric random variables with the same parameter p and let $q = 1 - p$, that is,

$$P(X_1 = r) = pq^{r-1}; \quad r = 1, 2, \dots$$

Denote $S_k = X_1 + X_2 + \dots + X_k$.

a. (10) Find $P(S_k = n)$ for $1 \leq k \leq n$.

Solution: since sums of independent geometric random variables are negative binomial, we have $S_k \sim \text{NegBin}(k, p)$. We have

$$P(S_k = n) = \binom{n-1}{k-1} p^k q^{n-k}.$$

b. (10) For each $n \geq 1$, compute

$$f_n = P(S_k = n \text{ for some } k = 1, 2, \dots, n).$$

Solution: we have

$$f_n = P(\cup_{k=1}^n \{S_k = n\}).$$

Observe that the events in the union are disjoint, so we get

$$f_n = \sum_{k=1}^n P(S_k = n).$$

Using the result from the first part and the binomial formula, we get

$$f_n = \sum_{k=1}^n \binom{n-1}{k-1} p^k q^{n-k} = p.$$

3. (20) Let the vector (U, X, Y) have the density

$$f(u, x, y) = \frac{x|y|}{2\pi\sqrt{u^3(1-u)^3}} e^{-\frac{x^2}{2u}} e^{-\frac{y^2}{2(1-u)}}$$

for $u \in (0, 1)$, $x > 0$ and $y \in \mathbb{R}$, and zero elsewhere. Define

$$W = \frac{X}{\sqrt{U}} \quad \text{and} \quad Z = \frac{Y}{\sqrt{1-U}}.$$

a. (10) Find the density of the vector (U, W, Z) . Are the random variables U , W and Z independent?

Solution: define

$$\Phi(u, x, y) = \left(u, \frac{x}{\sqrt{u}}, \frac{y}{\sqrt{1-u}} \right)$$

and observe that the map Φ takes $(0, 1) \times (0, \infty) \times \mathbb{R}$ bijectively onto itself. We have

$$\Phi^{-1}(u, w, z) = (u, w\sqrt{u}, z\sqrt{1-u}),$$

which implies $J_{\Phi^{-1}}(u, w, z) = \sqrt{u(1-u)}$. The transformation formula gives

$$f_{U,W,Z}(u, w, z) = \frac{\sqrt{u(1-u)} w|z|}{2\pi\sqrt{u^3(1-u)^3}} e^{-\frac{w^2}{2}} e^{-\frac{z^2}{2}} \cdot \sqrt{u(1-u)},$$

which simplifies to

$$f_{U,W,Z}(u, w, z) = \frac{w|z|}{2\pi\sqrt{u(1-u)}} e^{-\frac{w^2}{2}} e^{-\frac{z^2}{2}}.$$

We infer that U , W and Z are independent.

b. (10) Find the density of $(U, Y) = (U, Z\sqrt{1-U})$, and compute the density of Y .

Hint: when computing the marginal density, use the new variable

$$\frac{\sqrt{u}}{\sqrt{1-u}} = v.$$

Solution: noting that $\int_0^\infty w e^{-w^2/2} dw = 1$, we find that the vector (U, Z) has density

$$f_{U,Z}(u, z) = \frac{|z|}{2\pi\sqrt{u(1-u)}} e^{-\frac{z^2}{2}}.$$

Taking $\Phi(u, z) = (u, \sqrt{1-u} \cdot z)$, the transformation formula gives

$$f_{U,Y}(u, y) = f_{U,Z}(u, y/\sqrt{1-u}) \cdot \frac{1}{\sqrt{1-u}}.$$

Combining both equalities we get

$$f_{U,Y}(u, y) = \frac{|y|}{2\pi\sqrt{u(1-u)^3}} e^{-\frac{y^2}{2(1-u)}}.$$

The density of Y is the marginal density. Integrate the expression with respect to u . Introducing the new variable

$$\frac{\sqrt{u}}{\sqrt{1-u}} = v$$

we get

$$\frac{du}{2\sqrt{u(1-u)^3}} = dv, \quad \text{and} \quad \frac{1}{1-u} = 1 + v^2.$$

Finally,

$$\begin{aligned} f_Y(y) &= \frac{|y|}{\pi} \int_0^\infty e^{-\frac{y^2(1+v^2)}{2}} dv \\ &= \frac{\sqrt{2}|y|}{\sqrt{\pi}} e^{-\frac{y^2}{2}} \cdot \frac{1}{\sqrt{2\pi}} \int_0^\infty e^{-\frac{y^2 v^2}{2}} dv \\ &= \frac{\sqrt{2}|y|}{\sqrt{\pi}} e^{-\frac{y^2}{2}} \cdot \frac{1}{2|y|} \\ &= \frac{1}{\sqrt{2\pi}} e^{-\frac{y^2}{2}}. \end{aligned}$$

4. (20) In a sequence of independent tosses of a fair coin let X be the number of tosses until the first appearance of the pattern HH, and Y the number of tosses until the second appearance of the pattern HH. Examples:

$$\begin{array}{ll} \text{HTTHHTTTHTHH} & X = 5, \quad Y = 12 \\ \text{HTTHTTHTTHHH} & X = 11, \quad Y = 12 \end{array}$$

a. (10) Find $E(X)$.

Solution: define $B_1 = \{\text{first toss is a T}\}$, $B_2 = \{\text{first two tosses are HT}\}$, and $B_3 = \{\text{first two tosses are HH}\}$. We have

$$E(X|B_1) = 1 + E(X), \quad E(X|B_2) = 2 + E(X) \quad \text{and} \quad E(X|B_3) = 2.$$

The formula for total expectation gives

$$E(X) = \frac{1}{2}(1 + E(X)) + \frac{1}{4}(2 + E(X)) + \frac{1}{4} \cdot 2.$$

Solving the linear equation gives $E(X) = 6$.

b. (10) Find $E(Y - X)$.

Solution: for $k = 2, 3, \dots$ define

$$B_k = \{X = k, \text{ the } (k + 1)\text{-th toss is H}\}$$

and

$$C_k = \{X = k, \text{ the } (k + 1)\text{-th toss is T}\}.$$

We have

$$E(Y - X|B_k) = 1 \quad \text{and} \quad E(Y - X|C_k) = 1 + E(X).$$

Noting that the events $B_2, B_3, \dots, C_2, C_3, \dots$ form a partition, the formula for total expectation gives

$$E(Y - X) = \sum_{k=2}^{\infty} P(B_k) + \sum_{k=2}^{\infty} (1 + E(X))P(C_k).$$

Since $P(B_k) = P(C_k)$ and since these events form a partition, we have $\sum_{k=2}^{\infty} P(B_k) = \sum_{k=2}^{\infty} P(C_k) = \frac{1}{2}$. As a result, we conclude that

$$E(Y - X) = 1 + \frac{1}{2}E(X) = 4.$$

5. (20) Let X and Y be independent, non-negative, integer valued random variables with the same distribution. Assume that for $k \geq 1$ we have

$$P(X = k) = \frac{1}{4} P(X + Y = k - 1).$$

Let $G(s)$ be the generating function of X and Y .

a. (10) Find an equation that is satisfied by $G(s)$.

Solution: multiply both sides of the above relation by s^k and sum over $k \geq 1$. Denoting $P(X = 0) = p$, we get

$$\sum_{k=1}^{\infty} P(X = k) s^k = G_X(s) - p$$

and

$$\sum_{k=1}^{\infty} \frac{1}{4} P(X + Y = k - 1) s^k = \frac{s}{4} G_{X+Y}(s).$$

Since X and Y have the same distribution, we have $G_{X+Y}(s) = G(s)^2$. The desired equation is

$$G(s) - p = \frac{s}{4} G(s)^2.$$

b. (10) Find the distribution of X .

Hint: first $G(1) = 1$, and by Newton's expansion we have that for $|x| < 1$

$$\sqrt{1-x} = \sum_{k=0}^{\infty} (-1)^k \binom{1/2}{k} x^k.$$

Solution: since $G(1) = 1$, the equation from the first part implies

$$1 - p = \frac{1}{4}.$$

Solving for $G(s)$ we get

$$G(s) = \frac{2 \left(1 \pm \sqrt{1 - \frac{3s}{4}} \right)}{s}.$$

The coefficients of a generating function must be non-negative. Since $(-1)^k \binom{1/2}{k} < 0$ for all $k = 1, 2, 3, \dots$, we have to choose the negative sign for the root. Expanding into a power series we get

$$G(s) = \sum_{k=1}^{\infty} 2 \binom{1/2}{k} (-1)^{k-1} \frac{3^k s^{k-1}}{4^k}.$$

Finally,

$$P(X = k) = 2 \binom{1/2}{k+1} (-1)^k \left(\frac{3}{4} \right)^{k+1}.$$

6. (20) In 1999 the patrons of HIT Casinos played the game *Colore* 400,000 times. The probability of winning in the game is $p = 0.00198079$.

- a. (10) The number of winning games in the 440,000 games is like a sum of independent equally distributed indicators X_i with $P(X_i = 1) = p$ and $P(X_i = 0) = 1 - p$. What, approximately, is the probability that there are 920 or more winning games.

Solution: using the central limit theorem with the continuity correction and $n = 440,000$ we have

$$\begin{aligned} P(S_n \geq 920) &= P(S_n > 919.5) = P\left(\frac{S_n - E(S_n)}{\sqrt{\text{var}(S_n)}} \geq \frac{919.5 - E(S_n)}{\sqrt{\text{var}(S_n)}}\right) \\ &\approx P\left(Z \geq \frac{919.5 - E(S_n)}{\sqrt{\text{var}(S_n)}}\right), \end{aligned}$$

where $Z \sim N(0, 1)$. Denoting $q = 1 - p$, we have $E(S_n) = np = 871.55$ and $\sqrt{\text{var}(S_n)} = \sqrt{npq} = 29.49$ so

$$\frac{919.5 - E(S_n)}{\sqrt{\text{var}(S_n)}} \doteq \frac{919.5 - 871.55}{29.49} \doteq 1.63.$$

The normal table gives

$$P(S_n \geq 920) \approx 0.052.$$

Remark. Using exact binomial probabilities, we can compute a more accurate value. Within displayed accuracy, we have $P(S_n \geq 920) \doteq 0.5295$.

- b. (10) Suppose the payout for a winning game is $x > 0$. If a patron stakes 1 unit and wins, she gets the stake back along with additional x units. If she loses the game, she loses the stake. Find the largest x such that after 440,000 games the Casino has a loss with probability at most 0.01.

Solution: in every game the Casino either wins one unit or loses x units. The gain is like the sum of 440,000 independent equally distributed random variables Y_i with $P(Y_i = 1) = q$ and $P(Y_i = -x) = p$. Letting $W_n = Y_1 + Y_2 + \dots + Y_n$, the largest x has to approximately satisfy the equation

$$P(W_n < 0) = 0.01.$$

We compute $E(Y_i) = q - px$, $\text{var}(Y_i) = pq(x + 1)^2$ and $\text{var}(W_n) = npq(x + 1)^2$. By the central limit theorem we have

$$\begin{aligned} P(W_n < 0) &= P\left(\frac{W_n - E(W_n)}{\sqrt{\text{var}(W_n)}} < -\frac{E(W_n)}{\sqrt{\text{var}(W_n)}}\right) \\ &= P\left(Z < -\frac{(q - px)\sqrt{n}}{(x + 1)\sqrt{pq}}\right) \\ &\approx 0.01. \end{aligned}$$

From the normal table we infer that

$$-\frac{(q - px)\sqrt{n}}{(x + 1)\sqrt{pq}} \approx -2.33,$$

Solving for x we get

$$x \approx \frac{q\sqrt{n} - 2.33\sqrt{pq}}{p\sqrt{n} + 2.33\sqrt{pq}} \doteq 466.95.$$

Remark. Exact binomial probabilities again yield more accurate value. Keeping the notation from the preceding part, the gain of the Casino equals $-S_n x + (n - S_n)$. Thus, the Casino has a loss if and only if $S_n > \frac{n}{x+1}$. From exact binomial probabilities, it follows that $P(S_n > 940) > 0.01$ and $P(S_n > 941) < 0.01$. Therefore, $P(S_n > y) \leq 0.01$ if and only if $y \geq 941$. Hence the maximum value of x equals

$$x = \frac{n}{941} - 1 \doteq 466.5877.$$